

# Benchmarking and movement preservation. Evidences from real-life and simulated series

Tommaso Di Fonzo and Marco Marini

**Abstract** The benchmarking problem arises when time series data for the same target variable are measured at different frequencies with different level of accuracy, and there is the need to remove discrepancies between annual benchmarks and corresponding sums of the sub-annual values. Two widely used benchmarking procedures are the Modified Denton *Proportional First Differences* (PFD) and the Causey and Trager *Growth Rates Preservation* (GRP) techniques. The PFD procedure, which looks for benchmarked estimates aimed at minimizing the sum of squared proportional differences between the target and the unbenchmarked values, involves simple matrix operations. The GRP technique is a non-linear procedure based on a ‘true’ movement preservation principle, according to which the sum of squared differences between the growth rates of the target and of the unbenchmarked series is minimized. In the literature it is often claimed that the PFD procedure produces results very close to those obtained through the GRP procedure. In this paper we study the conditions under which this result holds, by looking at an artificial and a real-life economic series, and by means of a simulation exercise.

**Key words:** Time series analysis, Combining data from different sources, National economic and social accounts, Benchmarking, Movement preservation, Modified Denton PFD, Causey and Trager GRP

## 1 Introduction

Benchmarking monthly and quarterly series to annual series is a common practice in many National Statistical Institutes. The benchmarking problem arises when

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time series data for the same target variable are measured at different frequencies with different level of accuracy, and there is the need to remove discrepancies between annual benchmarks and corresponding sums of the sub-annual values. The most widely used benchmarking procedures are the Modified Denton *Proportional First Differences* (PFD) technique (Denton, 1971; Cholette, 1984), and the Causey and Trager (1981) *Growth Rates Preservation* (GRP) procedure (see also Bozik and Otto, 1988). The PFD procedure looks for benchmarked estimates aimed at minimizing the sum of squared proportional differences between the target and the unbenchmarked values, and is characterized by an explicit benchmarking formula involving simple matrix operations. The GRP technique is a non-linear procedure based on a ‘true’ movement preservation principle, according to which the sum of squared differences between the growth rates of the target and of the unbenchmarked series is minimized. As in the literature (e.g. Bloem *et al.*, 2001) it is often claimed that the PFD procedure produces results very close to those obtained through the GRP procedure, in this paper we study the conditions under which this result holds. We do that by showing how the two procedures work in practice, by looking at an artificial and a real-life economic series. Then a simulation exercise is performed in order to appreciate the impact on the benchmarked series of the variance of the observational error and of possible ‘steps’ in the annual benchmarks.

The paper is organized as follows. In section 2 the two benchmarking procedures are described, and the way they take into account a ‘movement preservation principle’ is discussed. In section 3 the artificial time series of Denton (1971) and a quarterly preliminary series of the EU Quarterly Sector Accounts (Di Fonzo and Marini, 2009) are benchmarked to their annual counterparts, using both Modified Denton PFD and Causey and Trager GRP benchmarking procedures, and the results are discussed. In section 4 we design a simulation exercise in order to analyze the distinctive features of the two procedures.

## 2 Two benchmarking procedures

Let  $Y_T$ ,  $T = 1, \dots, N$ , and  $p_t$ ,  $t = 1, \dots, n$ , be, respectively, the (say annual) totals and the (say quarterly) preliminary values of an unknown quarterly target variable  $y_t$ . The preliminary values being not in line with the annual benchmarks, i.e.  $\sum_{t \in T} p_t \neq Y_T$ ,  $T = 1, \dots, N$ , we look for benchmarked estimates  $y_t^b$  such that  $\sum_{t \in T} y_t^b = Y_T$ . As Bozik and Otto (1988, p. 2) stress, “Just forcing a series to sum to its benchmark totals does not make a unique benchmark series”. Some characteristic of the original series  $p_t$  should be considered in addition, in order to get benchmarked estimates ‘as close as possible’ to the preliminary values. In an economic time series framework, the preservation of the temporal dynamics (however defined) of the preliminary series is often a major interest of the practitioner. Thus in what follows we consider two procedures designed to preserve at the best the movement of the series  $p_t$ .

Denton (1971) proposed a benchmarking procedure grounded on the *Proportionate First Differences* between the target and the original series. Cholette (1984) slightly modified the result of Denton, in order to correctly deal with the starting conditions of the problem. The PFD benchmarked estimates are thus obtained as the solution to the constrained quadratic minimization problem

$$\min_{y_t} \sum_{t=2}^n \left( \frac{y_t}{p_t} - \frac{y_{t-1}}{p_{t-1}} \right)^2 \quad \text{subject to} \quad \sum_{t \in T} y_t = Y_T, \quad T = 1, \dots, N. \quad (1)$$

In matrix notation, denoting  $\mathbf{p}$  and  $\mathbf{Y}$  the  $(n \times 1)$  and  $(N \times 1)$ , respectively, vectors of preliminary and benchmark values, the PFD benchmarked series is contained in the  $(n \times 1)$  vector  $\mathbf{y}^{PFD}$  solution of the linear system (Di Fonzo and Marini, 2009)

$$\begin{bmatrix} \mathbf{Q} & \mathbf{C}' \\ \mathbf{C} & \mathbf{0} \end{bmatrix} \begin{bmatrix} \mathbf{y}^{PFD} \\ \lambda \end{bmatrix} = \begin{bmatrix} \mathbf{Q}\mathbf{p} \\ \mathbf{Y} \end{bmatrix},$$

where  $\lambda$  is a  $(N \times 1)$  vector of Lagrange multipliers,  $\mathbf{Q} = \mathbf{P}^{-1} \Delta_n' \Delta_n \mathbf{P}^{-1}$ ,  $\mathbf{P} = \text{diag}(\mathbf{p})$ ,  $\mathbf{C}$  is a  $(N \times n)$  temporal aggregation matrix converting quarterly values in their annual sums, and  $\Delta_n$  is the  $((n-1) \times n)$  first differences matrix.

Notice that  $\mathbf{P}^{-1}\mathbf{p} = \mathbf{1}_n$  and  $\Delta_n \mathbf{1}_n = \mathbf{0}$ , so that it is  $\mathbf{Q}\mathbf{p} = \mathbf{0}$ . In addition, the coefficient matrix of the system can be factorized as

$$\begin{bmatrix} \mathbf{Q} & \mathbf{C}' \\ \mathbf{C} & \mathbf{0} \end{bmatrix} = \begin{bmatrix} \mathbf{P}^{-1} & \mathbf{0} \\ \mathbf{0} & \mathbf{I}_N \end{bmatrix} \begin{bmatrix} \Delta_n' \Delta_n & \mathbf{P}\mathbf{C}' \\ \mathbf{C}\mathbf{P} & \mathbf{0} \end{bmatrix} \begin{bmatrix} \mathbf{P}^{-1} & \mathbf{0} \\ \mathbf{0} & \mathbf{I}_N \end{bmatrix},$$

whose inverse can be expressed as

$$\begin{bmatrix} \mathbf{Q} & \mathbf{C}' \\ \mathbf{C} & \mathbf{0} \end{bmatrix}^{-1} = \begin{bmatrix} \mathbf{P} & \mathbf{0} \\ \mathbf{0} & \mathbf{I}_N \end{bmatrix} \begin{bmatrix} \Delta_n' \Delta_n & \mathbf{P}\mathbf{C}' \\ \mathbf{C}\mathbf{P} & \mathbf{0} \end{bmatrix}^{-1} \begin{bmatrix} \mathbf{P} & \mathbf{0} \\ \mathbf{0} & \mathbf{I}_N \end{bmatrix}.$$

The solution of the above linear system can thus be written as

$$\begin{bmatrix} \mathbf{y}^{PFD} \\ \lambda \end{bmatrix} = \begin{bmatrix} \mathbf{P} & \mathbf{0} \\ \mathbf{0} & \mathbf{I}_N \end{bmatrix} \begin{bmatrix} \Delta_n' \Delta_n & \mathbf{P}\mathbf{C}' \\ \mathbf{C}\mathbf{P} & \mathbf{0} \end{bmatrix}^{-1} \begin{bmatrix} \mathbf{P} & \mathbf{0} \\ \mathbf{0} & \mathbf{I}_N \end{bmatrix} \begin{bmatrix} \mathbf{0} \\ \mathbf{Y} \end{bmatrix},$$

that is

$$\begin{bmatrix} \mathbf{y}^{PFD} \\ \lambda \end{bmatrix} = \begin{bmatrix} \mathbf{P} & \mathbf{0} \\ \mathbf{0} & \mathbf{I}_N \end{bmatrix} \begin{bmatrix} \Delta_n' \Delta_n & \mathbf{P}\mathbf{C}' \\ \mathbf{C}\mathbf{P} & \mathbf{0} \end{bmatrix}^{-1} \begin{bmatrix} \mathbf{0} \\ \mathbf{Y} \end{bmatrix}. \quad (2)$$

Causey and Trager (1981; see also Monsour and Trager, 1979, and Trager, 1982) consider a different quadratic minimization problem, in which the criterion to be minimized is explicitly related to the growth rate, which is a natural measure of the movement of a time series:

$$\min_{y_t} \sum_{t=2}^n \left( \frac{y_t}{y_{t-1}} - \frac{p_t}{p_{t-1}} \right)^2 \quad \text{subject to} \quad \sum_{t \in T} y_t = Y_T, \quad T = 1, \dots, N. \quad (3)$$

Looking at the criterion to be minimized in (3), it clearly appears that, differently from (1), it is grounded on an “ideal” movement preservation principle, “formulated as an explicit preservation of the period-to-period rate of change” of the preliminary series (Bloem *et al.*, 2001, p. 100).

It should be noted that, while problem (1) has linear first-order conditions for a minimum, and thus gives rise to an explicit solution as shown in (2), the minimization problem in (3) is inherently non-linear. Trager (1982; see Bozik and Otto, 1988) suggests to use a technique based on the steepest descent method, using  $\mathbf{y}^{PFD}$  as starting values, in order to calculate the benchmarked estimates  $y_t^{GRP}$ ,  $t = 1, \dots, n$ , solution to problem (3). We employ the Sequential Quadratic Programming (SQP) method available in the Optimization Toolbox<sup>TM</sup> of MATLAB<sup>®</sup> (version 2009b). It consists of an iterative procedure that splits the problem into easier subproblems to be solved by standard (quadratic) nonlinear programming methods. In each iteration an approximation of the Hessian of the Lagrangian function is calculated using a quasi-Newton updating method, which enables to make informed decisions regarding directions of search and step length.

It is interesting to go deep into the relationship between the criteria optimized by the two alternative procedures. Let

$$C_{PFD} = \sum_{t=2}^n \left( \frac{y_t}{p_t} - \frac{y_{t-1}}{p_{t-1}} \right)^2 \quad \text{and} \quad C_{GRP} = \sum_{t=2}^n \left( \frac{y_t}{y_{t-1}} - \frac{p_t}{p_{t-1}} \right)^2$$

be the objective functions of the PFD and GRP benchmarking procedures, respectively. After a bit of algebra, we can write (U.S. Bureau of the Census, 2009, p.96):

$$C_{PFD} = \sum_{t=2}^n \left[ \frac{y_{t-1}}{p_t} \left( \frac{y_t}{y_{t-1}} - \frac{p_t}{p_{t-1}} \right) \right]^2 \quad (4)$$

Expression (4) makes clear the relationship between  $C_{PFD}$  and  $C_{GRP}$ . The term in parentheses is the difference between the growth rates of the target and the preliminary series, namely the addendum of  $C_{GRP}$ . In  $C_{PFD}$  these terms are weighted by the ratio between the target series at  $t - 1$  and the preliminary series at  $t$ . When these ratios are relatively stable over time, which is the case when the ‘benchmark-to-indicator ratio’  $\frac{\dot{Y}_T}{\sum_{t \in T} p_t}$ ,  $T = 1, \dots, N$  (Bloem *et al.*, 2001), is a smooth series,  $C_{PFD}$  and  $C_{GRP}$  are very close to each other. On the contrary, when the ratios  $(y_{t-1}/p_t)$  behave differently each term in the summation is over-(under-)weighted according to the specific relationship between target and preliminary series in that period. For example, sudden breaks in the movements of  $y_{t-1}/p_t$  might arise in case of large differences between the annual benchmarks and the annually aggregated preliminary series. Keeping in mind this relationship, we move to investigate on the differences between the PFD and the GRP benchmarking solutions in simulated and real-life cases.

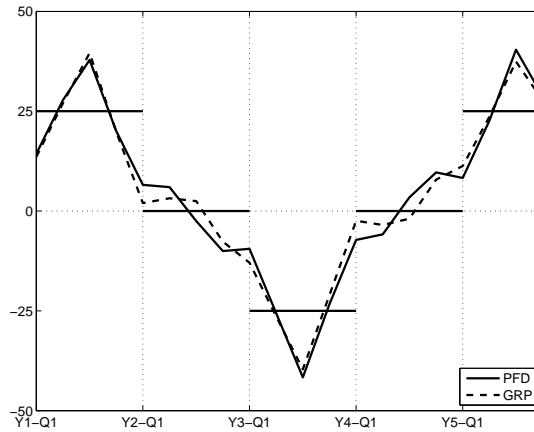
### 3 Evidences from artificial and real time series

In this section we apply both the PFD and GRP benchmarking procedures to two illustrative examples, in order to show to what extent the former solution can be used effectively to approximate the “ideal” movement preservation criterion based on growth rates. The results will be evaluated in terms of the ratio of the ‘Growth Rate Preservation’ criteria computed from the two solutions, that is

$$r = \frac{\sum_{t=2}^n \left( \frac{y_t^{GRP}}{y_{t-1}^{GRP}} - \frac{p_t}{p_{t-1}} \right)^2}{\sum_{t=2}^n \left( \frac{y_t^{PFD}}{y_{t-1}^{PFD}} - \frac{p_t}{p_{t-1}} \right)^2}. \quad (5)$$

Obviously, we expect the GRP technique always reaches a lower value of the chosen criterion than PFD, and thus the ratio (5) is never larger than 1.

The first example we consider is the artificial preliminary series used in the seminal paper of Denton (1971). It consists of a five-year artificial quarterly series, with a fixed seasonal pattern invariant from year to year. The values are 50, 100, 150 and 100 in the four quarters, for a total yearly amount of 400. The annual benchmarks are assumed to be 500, 400, 300, 400 and 500 in the five successive years. The corresponding discrepancies (i.e., the differences between the known benchmarks and the annual sums of the preliminary series) are therefore 100, 0, -100, 0 and 100, respectively. The benchmarked values are shown in table 2: as expected, the minimum  $C_{GRP}$  is achieved by the GRP procedure (441.2 against 1442.8 of PFD, with  $r=0.306$ ). Figure 1 shows the adjustments to the levels of the original series in the two cases. The horizontal lines in each year denote the (average) annual discrepancy to be distributed.



**Fig. 1** Adjustments to the artificial series produced by the PFD and GRP procedures

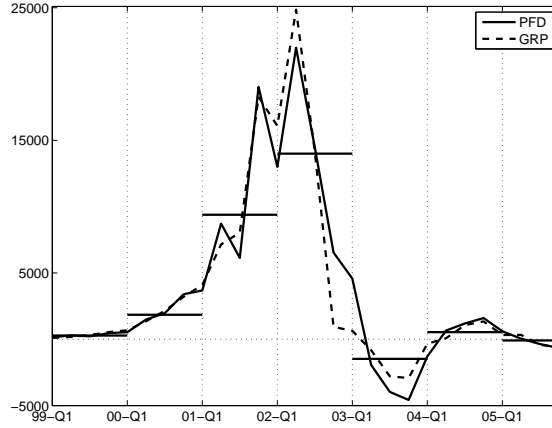
**Table 1** Adjustments to the artificial series produced by the PFD and GRP procedures

Year	Qtr	$p_t$	$y_t^{PFD}$	$y_t^{GRP}$	$y_t^{PFD} - p_t$	$y_t^{GRP} - p_t$
1	1	50	64.3	63.6	14.3	13.6
	2	100	127.8	127.0	27.8	27.0
	3	150	187.8	189.6	37.8	39.6
	4	100	120.0	119.8	20.0	19.8
2	1	50	56.6	52.0	6.6	2.0
	2	100	106.0	103.2	6.0	3.2
	3	150	147.5	152.5	-2.5	2.5
	4	100	90.0	92.3	-10.0	-7.7
3	1	50	40.5	37.1	-9.5	-12.9
	2	100	74.4	73.6	-25.6	-26.4
	3	150	108.3	110.3	-41.7	-39.7
	4	100	76.7	79.0	-23.3	-21.0
4	1	50	42.8	47.6	-7.2	-2.4
	2	100	94.1	96.5	-5.9	-3.5
	3	150	153.4	148.1	3.4	-1.9
	4	100	109.7	107.9	9.7	7.9
5	1	50	58.3	61.3	8.3	11.3
	2	100	122.6	123.6	22.6	23.6
	3	150	190.4	187.4	40.4	37.4
	4	100	128.7	127.7	28.7	27.7

The second example is a real-life economic series coming from the European Quarterly Sector Accounts (EU-QSA). The EU-QSA system has been dealt with by Di Fonzo and Marini (2009) in a *reconciliation* exercise, where several time series have to be adjusted in order to be in line with both temporal and contemporaneous known aggregates (Dagum and Cholette, 2006). In this paper we consider the series 'Other Property Income' of the Financial Corporation sector, showing a considerable amount of temporal discrepancies. Figure 2 shows the large discrepancy in 2002, when the original series accounts for just 65% of the annual target. From 2003 onwards the discrepancies are much more contained. This is a typical practical situation where the preservation of the original growth rates can be better guaranteed by the GRP procedure ( $r = 0.355$ ). The quarterly adjustments in the two cases are also displayed in figure 2. The differences are large in the years with large discrepancies (2001-2002), but they are also remarkable in 2003, when the discrepancy is limited. In this case the smoother distribution produced by the GRP procedure is clearly visible.

#### 4 The simulation exercise

By means of this experiment we wish to shed light on the conditions under which the PFD benchmarking procedure produces results 'close' to the GRP technique in terms of differences between the growth rates of the benchmarked and preliminary series. We consider quarterly series covering a period of 7 years ( $n = 28$ ).



**Fig. 2** Adjustments to the real-life series produced by the PFD and GRP procedures

Let  $\theta_t = \theta_{t-1} + \varepsilon_t$  be a random walk process, where  $\varepsilon_t$  is a Gaussian white noise with unit variance ( $\sigma_\varepsilon^2 = 1$ ) and  $\theta_0 = \varepsilon_0$ . The target series of the exercise,  $y_t$ , is derived as

$$y_t = \theta_t^* + \mu_t, \quad t = 1, \dots, n$$

with  $\theta_t^* = 100 + \theta_t$ , where the constant term 100 is large enough to prevent negative values, and  $\mu_t$  is given by

$$\mu_t = \begin{cases} \mu & t = 9, \dots, 16 \\ -\mu & t = 17, \dots, 24 \\ 0 & \text{elsewhere} \end{cases}$$

The preliminary series  $p_t$  is related to  $y_t$  as follows:

$$p_t = \theta_t^* + e_t, \quad t = 1, \dots, n$$

where  $e_t$  is a Gaussian white noise with variance  $\sigma_e^2$ . It is clear that preliminary and target series are different for the effects of  $\mu_t$  and  $e_t$ . The former is introduced in the model for  $y_t$  in order to simulate yearly biases of the preliminary series. The first control parameter of the experiment is thus  $\mu$ . When  $\mu > 0$ , the target series contains a positive drift from  $p_t$  in years 3 and 4, followed by a negative step (of the same amount) in years 5 and 6. We set  $\mu = 0, 15, 30, 45, 60$ . The second control parameter is  $\sigma_e$ , the standard deviation of the innovation process  $e_t$ . The larger this parameter is, the larger the observational error in the preliminary series will be. We set  $\sigma_e = 5, 10, 15, 20, 25$ .

We drew two sets of 1,000  $n$ -dimensional vectors as  $N(0, 1)$ . One set is used to simulate  $\varepsilon_t$ ; the other is used to derive the innovation  $e_t$  according to the 5 levels of  $\sigma_e$ . By using the 5 values of  $\mu$ , we achieved 1,000 experiments for each of the 25

combinations. For each combination, we computed summary statistics on the ratios  $r$  obtained over the 1,000 experiments.

Table 2 shows the median<sup>1</sup> value of  $r$  under different values of  $\sigma_e$  (rows) and  $\mu$  (columns). As expected, the median ratio is always smaller than 1. However, the PFD procedure provides very similar results to GRP when discrepancies are small and unsystematic (median ratio  $\geq 0.9$  when  $\sigma_e \leq 10$  and  $\mu \leq 15$ ). The reduction is stronger as both  $\sigma_e$  and  $\mu$  increase. This implies that, when the uncertainty of the preliminary series and/or its bias with respect to the target variable are high, the GRP procedure is likely to provide better results (in terms of growth rates) than PFD.

**Table 2** Median value of  $r$  for different values of  $\sigma_e$  and  $\mu$  (across 1,000 experiments)

$\sigma_e$	$\mu$				
	0	15	30	45	60
5	0.984	0.969	0.919	0.816	0.583
10	0.936	0.922	0.872	0.770	0.542
15	0.859	0.846	0.799	0.700	0.489
20	0.759	0.744	0.705	0.610	0.429
25	0.641	0.624	0.592	0.507	0.364

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<sup>1</sup> The median is more representative than the mean in the case of atypical values. We also calculated mean, standard deviation, minimum, maximum and range of the ratios, available on request from the authors.